

What Does Performance in Graduate School Predict? Graduate Economics Education and Student Outcomes

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Economists devote considerable effort to graduate student education, but have conducted relatively little research on the determinants of student performance while in graduate school or placement in the job market. Do graduate students who do well in core microeconomics (Micro) courses also do well in core macroeconomics (Macro) and econometrics (Metrics) courses? Are students who achieve higher grades in their first-year core classes or general exams more likely to complete their Ph.D. and to obtain higher ranked positions in the job market? In an attempt to answer these questions, we assembled a rich data set on 1,030 economics graduate students who enrolled at Harvard, the Massachusetts Institute of Technology (MIT), Princeton, Stanford or the University of Chicago in the 1990s. These schools were selected because they have the five highest ranked economics Ph.D. programs according to the National Research Council in 1993.

Our results indicate that students' grades in required core courses are highly correlated across subjects. The Ph.D. admissions committee's evaluation of a student predicts first-year grades and Ph.D. completion, but not job placement. First-year performance is a strong predictor of Ph.D. completion. Most importantly, we find that first-year Micro and Macro grades are statistically significant predictors of student job placement, even conditional on Ph.D. completion. Conditional on first-year grades, GRE scores, foreign citizenship, sex, and having a prior Masters degree do not predict job placement. Students who attended elite undergraduate universities and liberal arts colleges are more likely to be placed in top ranked academic jobs.

I. Data and Descriptive Statistics

After receiving permission from each university's Institutional Review Board, in the summer of 2006 we attempted to gather information on every student who entered the Ph.D. program in economics at Chicago, Harvard, MIT, Princeton, or Stanford from 1990 to 1999. Specifically, we sought data on GRE scores, admissions rank, first-year course or general exam grades, Ph.D. completion status, initial job placement and other variables from the schools' economics department administrators and archives.

Some departments were better at record keeping than others. Data are available for Princeton (n=217) and Stanford (n=259) for students who enrolled from 1990-99. Chicago (n=230) and Harvard (n=177) provided data on enrollees from 1994-99. MIT provided data on students who completed their degree from 1990-99; we restrict the MIT sample to 147 students who enrolled between 1990 and 1995 *and* completed their Ph.D. We drop MIT when we analyze the determinants of degree completion. The rank the Ph.D. admissions committee assigned the student is available for Harvard, Princeton, and Stanford.

Harvard, Stanford and Princeton (before 1996) relied on a general exam grade for their requirements in Micro and Macro. Chicago relied on both course grades and general exams for Ph.D. advancement. MIT and Princeton (after 1995) used course grades for degree requirements. We computed the average first-year course grade in Micro and in Macro for Chicago, MIT, and Princeton (after 1995). Stanford gave a first-year general in Econometrics while the other schools use course grades in Econometrics. To convert grades into comparable units across schools, we rescaled all course and general exam grades into percentile ranks within year and school.¹

Some schools provided information on students' initial job placements while others did not. In most cases, we were able to identify job placement even when the school could not provide it. Job placement is typically not available for those who did not complete their Ph.D. degree; for 93 percent of completers, we have initial job information. Rating the quality of job placement is inherently a subjective judgment, so we use multiple measures. First, we used Pantelis Kalaitzidakis, et al.'s (2003) ranking of the top 200 economics departments. This ranking, which is based on publications, was selected because it is relatively comprehensive, but we still needed to assign ranks to some positions ourselves.² Second, we created an indicator variable for placement at one of the twenty top-ranked economics departments according to Kalaitzidakis, et al. This group overlaps substantially with the NRC's top 20 departments. Third, we created a categorical variable that equals 1 for jobs with independent research opportunities, such as those at research universities and the Federal Reserve, 2 for primarily teaching positions such as those at many liberal arts colleges, 3 for directed research positions such as those at the Treasury Department and MDRC, and 4 for other positions as well as noncompleters. Finally, we created an indicator for academic jobs.

Means for the full sample are presented in column (1) of Table 1. To conceal the identity of the individual schools but still illustrate the range of variability across the five schools, in column (2) we report the mean for the school with the lowest average value and in column (3) the highest average value for each variable.

A quarter of the students who enrolled in these graduate programs are women.³ Under-represented minority students (African American, Hispanic or Native American) are only 3 percent of enrolled students. Half of the students attended a foreign

undergraduate institution, and 63 percent are foreign citizens. Twenty-two percent of enrollees graduated from a top 15 university as an undergraduate, based on the 2006 *U.S. News and World Report* ranking, and another 4 percent had attended one of the top five liberal arts colleges. The average student scored in the 93rd percentile on the Math GRE; three quarters of students scored above the 90th percentile. The average student scored in the 75th percentile of the Verbal GRE; the average for foreign citizens was the 66th percentile and the average for American citizens was the 90th percentile.

Overall, 75 percent of enrollees completed their Ph.D. degree; this figure ranged from 65 percent at one school to 90 percent at another.⁴ Twenty-seven percent of Ph.D. recipients from these five schools accepted jobs at “top 20” economics departments, while 17 percent landed a “top 10” job. More than half of graduates accepted an academic job, consistent with David Colander’s (2005) finding that most economics graduate students at top programs plan to pursue an academic job at some point.

II. Grades and Completion

For the sample as a whole, first-year grades (scaled as percentile ranks) are strongly positively correlated across subjects: 0.54 for Micro and Macro; 0.56 for Micro and Econometrics; and 0.56 for Macro and Metrics. These correlations are very similar in magnitude when computed separately for foreign and U.S. citizens. The Micro-Macro correlation ranged from a low of 0.32 in one school to a high of 0.73 in another. The strong correlations between grades in different subjects is consistent with the view that successful course (or test) performance in economics requires similar skills across fields, as well as with the view that the study habits and personal traits that lead to success in one subject also lead to success in others.

Table 2 presents Ordinary Least Squares (OLS) regression estimates in which the dependent variable is the student's percentile rank in first-year Micro, Macro or Econometrics. Explanatory variables are student background characteristics and admissions rank. The sample size drops nearly in half when we control for the admission rank as it is not available for Chicago and MIT. Several findings are noteworthy. The Analytical GRE score is a stronger predictor of grades in all three subjects than is the Math or Verbal GRE. When admissions rank is excluded from the model, the Math GRE has a statistically significant effect on Micro grades and the Verbal GRE has a statistically significant effect on Macro grades. The latter finding is partly a result of the fall in sample, as the Verbal GRE has a small and insignificant effect if the column (3) model is estimated with the column (4) sample. The Analytical GRE score has a sizable effect: an increase of 20 percentile points is associated with an increase in grades of 5 to 6 percentiles in all three subjects in the models that exclude the admissions rank. Admissions rank presumably reflects, in part, GRE scores.

Students who attended a foreign undergraduate institution perform significantly better in all three subjects. The first-year grades of female students are significantly lower than the grades of male students. Students who attended an elite undergraduate institution do not perform significantly differently than other students, with the possible exception of achieving a higher grade in Macro. Students who start graduate school with a masters degree in economics or finance do not perform better in Micro or Econometrics, but their Macro grade is 11 percentile points higher, all else equal. If we drop admissions rank, a prior masters degree is associated with an increase of 4.4 points in Micro ($p=.097$), 10.6 points in Macro ($p=.000$), and 3.1 points in Econometrics ($p=.36$). Lastly, the

admission rank is a strong predictor of performance: moving ahead 30 places in the ranking (from 50 to 20, say) is associated with an increase in grades of 7 percentile point in all three subjects, conditional on GREs and other factors.

We have estimated probit equations to examine the determinants of Ph.D. completion by 2006.⁵ The explanatory variables used in Table 2 as well as first-year grades were included as predictors. (MIT was excluded from this analysis because its sample omits dropouts.) These models indicate that first-year Micro, Macro and Metrics grades are strong predictors of Ph.D. completion. An increase in first-year Micro and Macro grades of 20 percentile points, for example, is associated with an increase in Ph.D. completion of 12 percentage points, which would cut the dropout rate nearly in half. One concern, however, is that grades have a mechanical relationship with completion as students who fail to pass their first-year courses or general exams are forced to leave the program. Even if we restrict the sample to those who scored above the 30th percentile on Micro and Macro – and are therefore unlikely to be forced out of the program – the effect of grades is about half as large but still statistically significant.

Conditional on grades, the only variable that significantly predicts completion in our sample is matriculation at a non-U.S. undergraduate institution. Those who completed their undergraduate education abroad have a 10 percentage point higher completion rate than American students. We do not find a gender difference in completion in our data. Students who attended an elite undergraduate college are more likely to complete their degree, but the effect is marginally significant. If grades are omitted, the admissions rank is a significant predictor of completion, with students who are ranked 30 places ahead being 4.5 percentage points more likely to complete a Ph.D.

III. Job Placement

Table 3 presents results of our job placement analysis. Columns 1-3 present OLS estimates where the dependent variable is the rank of each student's initial job. Columns 4-6 present probit estimates where the dependent variable equals 1 if the student was placed in a job in a top 20 position and 0 otherwise. The sample for the probit estimates consists of all enrollees, whereas the sample for columns 1-3 is mainly restricted to those who completed their degree. (We have job placements for 40 students who did not complete their degree who are included in columns 1-3.) In all of the models, a Chow test does not reject the null hypothesis that data can be pooled across the five schools.⁶

In columns (1) and (4) we predict job placement from variables that could be observed prior to the start of graduate school. Only the Analytical GRE score is statistically significant in predicting placement in a "top 20" economics department. Because this sample is conditional on acceptance to a top five economics department, the weak predictive power of GRE scores may be an artifact: students with low GRE scores who were admitted probably had other distinguishing characteristics that improved their job prospects. Krueger and Wu (2000) find that GREs are a significant predictor of job placement in a sample of all applicants to one top five economics department, but not of grades for matriculants. Students from top undergraduate institutions are substantially more likely to secure highly ranked positions after graduate school. The effect is large: in column (4), attending a leading liberal arts college is associated with an 18 percentage point increase in the chance of job placement at a "top 20" economics department.

Columns (2) and (5) add first-year grades and columns (3) and (6) add admission rank and a dummy for a prior masters degree. Micro and Macro grades are significant

even conditional on grades. Thus, it is something about their backgrounds that is not reflected in their first-year course performance that helps raise their job placements.

Table 4 presents results of using two additional approaches to rating jobs. In columns 1-3 we use an ordered categorical variable that groups students into independent research jobs, teaching positions, directed research, and other positions. In columns 4-6 we use a dichotomous variable that equals 1 if the student is initially placed in an academic job and 0 otherwise. The same qualitative conclusions emerge from these alternative measures of job placements with two noteworthy exceptions. First, students whose undergraduate education is from a non-U.S. institution are more likely to be placed in the independent research end of the job spectrum. Second, without controlling for grades, a smaller fraction of women than men are placed in jobs that allow for independent research after they complete graduate school.

IV. Conclusion

First-year grades in core required courses are a strong predictor of economics graduate students' job placements. The reasons for this relationship are unclear and a worthy topic of further research. One possibility is that mastering the content of first-year courses directly helps prepare students to be successful researchers.⁷ Other possibilities are that students who do well in their first-year courses gain self-confidence or create positive "first-impressions" with faculty members, putting them on a positive trajectory irrespective of the direct utility of the course content. Yet another interpretation is that the traits that enable a student to do well in the first-year of graduate school, such as high cognitive ability and diligence, are also important when it comes to writing a dissertation and searching for a job.

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Notes:

* Harvard University, Harvard University, Princeton University, University of Chicago, and MIT. Krueger is the corresponding author. Unfortunately, due to university confidentiality restrictions, we are not able to make the data used in this article publicly available. We thank Eleanor Choi, Elaine Liu, Yuhang Wang, Stephen Nei (Princeton), Peter Hoagland and Katherine Swan (MIT) for excellent research assistance, and Ed Freeland for preparing IRB submissions. We acknowledge partial research support from the NSF.

¹ For students who failed an exam or course and retook it, we used the last score available. For Chicago, however, we averaged the retake grade with the original grade because it was unclear which grade was the final one. Grades are not available for Stanford in 1990. Princeton only recorded pass/fail for Econometrics before 1996, so we only include Econometrics grades for Princeton for 1996-99.

² Following Krueger and Wu (2000), if a job at a particular school was not in the economics department, we added 5 ranks to the institution’s ranking. Jobs in the business sector were typically assigned a rank of 200 to 250, depending on whether they utilized research or teaching skills. Other jobs were assigned a rank that seemed commensurate with one of the departments ranked in Kalaitzidakis, *et al.* (2003).

³ Though not reported in the table, women also make up a quarter of the Ph.D. degree recipients in our sample.

⁴ MIT is excluded from these figures because we lack information on dropouts.

⁵ See Siegfried and Stock (2001) for an analysis of the time to degree for a broader sample of economics graduate students.

⁶ We also note that our conclusions are not qualitatively altered if we add unrestricted entry year dummies or school dummies.

⁷ Consistent with this interpretation, Stock and Hansen (2004) conclude that a majority of new economics Ph.D.s believe that “the emphasis given in their Ph.D. programs to many economic proficiencies was ‘about right’.”

**Table 1: Mean Characteristics of Economics Ph.D. Students Entering
Top 5 Departments from 1990 to 1999**

Characteristics	All	Lowest School	Highest School
Female	0.25	0.20	0.30
Under Represented Minority	0.03	0.01	0.06
Foreign Citizen	0.63	0.50	0.68
Graduated from Foreign Undergrad University	0.49	0.39	0.61
Graduated from Top 15 University	0.22	0.13	0.34
Graduated from Top 5 Liberal Arts College	0.04	0.00	0.06
Prior Masters Degree in Econ/Finance	0.24	0.16	0.30
GRE Quantitative (Percentile)	93	92	95
GRE Verbal (Percentile)	75	66	81
GRE Analytical (Percentile)	89	85	91
Admissions Rank	38	36	39
Ph.D. Completion Rate	0.75	0.65	0.90
Academic Job (1=yes)	0.45	0.30	0.56
Average Job Rank	74	61	86
Top 20 Jobs (Unconditional)	0.19	0.12	0.26
Top 20 Jobs (Conditional on Ph.D.)	0.27	0.20	0.31
Sample Size	1030		

Notes: Maximum sample size is 1,030. MIT is excluded from the sample used for the Ph.D. completion rate and the unconditional top 20 jobs because the MIT data excludes dropouts.

Table 3: Determinants of Initial Job Placements

Variable	Job Rank			Top 20		
	(1)	(2)	(3)	(4)	(5)	(6)
Micro Grade (percentile)	---	-0.309** (0.105)	-0.271* (0.141)	---	0.00313** (0.00060)	0.00321** (0.00084)
Macro Grade (percentile)	---	-0.231** (0.104)	-0.362** (0.140)	---	0.00193** (0.00059)	0.00207** (0.00081)
GRE Math (percentile)	0.049 (0.413)	0.144 (0.409)	0.363 (0.515)	0.00308 (0.00270)	0.00162 (0.00249)	0.00438 (0.00370)
GRE Verbal (percentile)	-0.124 (0.125)	-0.101 (0.125)	-0.061 (0.165)	0.00047 (0.00073)	0.00045 (0.00074)	0.00015 (0.00100)
GRE Analytical (percentile)	-0.283 (0.225)	-0.092 (0.227)	-0.070 (0.311)	0.00456** (0.00149)	0.00250* (0.00147)	0.00177 (0.00209)
Female	4.363 (5.798)	-1.084 (5.831)	-1.486 (7.901)	-0.01266 (0.03167)	0.04855 (0.03521)	0.03957 (0.04832)
Undergrad Top 15 University	-17.365** (7.153)	-16.929** (7.074)	-17.887* (10.035)	0.14816** (0.04603)	0.13367** (0.04649)	0.11515* (0.06656)
Undergrad Top 5 Liberal Art	-18.730 (12.611)	-21.282* (12.445)	-30.838* (15.900)	0.18424** (0.09076)	0.20652** (0.09218)	0.26646** (0.11745)
Undergrad Foreign University	-7.382 (6.760)	-1.802 (6.822)	-6.269 (9.804)	0.06118 (0.03868)	-0.00355 (0.03930)	0.06809 (0.05966)
Admissions Rank	---	---	-0.016 (0.152)	---	---	-0.00001 (0.00094)
Masters in Econ/Fin	---	---	-0.821 (8.362)	---	---	-0.04904 (0.04617)
R-squared	0.02	0.05	0.06	0.04	0.12	0.10
P-value for GRE Scores	0.2084	0.7272	0.8713	0.0006	0.1026	0.3758
R-square						
Observations	710	701	397	871	839	468

Notes: Columns 1-3 estimated by OLS. Coefficients in columns 4-6 reexpress probit coefficients as changes in probability. Students who never completed the Ph.D. are classified as not having top 20 jobs in the probit model. P-value for GRE scores is from an F-test that the 3 scores jointly equal zero.

* = p-value < .10 **=p-value<.05

Mean [standard deviation] of job rank is 74 [65]; mean top 20 is 0.21 [0.41]. The mean of top 20 differs from Table 1 because MIT is included in the probit sample but not in Table 1 top 20 (unconditional).

Table 4: Models for Type of Job Placement Category and Placement in an Academic Job

	Type of Initial Job Placement			Academic Position		
	(1)	(2)	(3)	(4)	(5)	(6)
Micro Grade Percentile	---	-0.0084** (0.0018)	-0.0084** (0.0025)	---	0.0031** (0.0008)	0.0030** (0.0010)
Macro Grade Percentile	---	-0.0078** (0.0018)	-0.0070** (0.0024)	---	0.0028** (0.0007)	0.0021** (0.0010)
GRE Math	-0.0006 (0.0069)	0.0027 (0.0071)	-0.0034 (0.0093)	0.0022 (0.0029)	0.0010 (0.0029)	0.0050 (0.0039)
GRE Verbal	-0.0023 (0.0021)	-0.0022 (0.0022)	-0.0022 (0.0031)	0.0008 (0.0009)	0.0009 (0.0009)	0.0006 (0.0012)
GRE Analytical	-0.0025 (0.0038)	0.0041 (0.0041)	0.0143** (0.0061)	0.0015 (0.0015)	-0.0009 (0.0016)	-0.0032 (0.0024)
Female	0.1654* (0.0975)	0.0070 (0.1028)	-0.1844 (0.1408)	-0.0315 (0.0398)	0.0298 (0.0423)	0.0814 (0.0563)
Undergrad Top 15 University	-0.3078** (0.1182)	-0.2567** (0.1232)	-0.1842 (0.1664)	0.1584** (0.0484)	0.1425** (0.0505)	0.1031 (0.0686)
Undergrad Top 5 Liberal Art University	-0.2200 (0.2172)	-0.2028 (0.2217)	-0.2583 (0.2746)	-0.0106 (0.0916)	-0.0139 (0.0935)	0.0059 (0.1139)
Undergrad Foreign University	-0.4128** (0.1116)	-0.2445** (0.1194)	-0.4025** (0.1714)	0.1141** (0.0454)	0.0506 (0.0488)	0.1128 (0.0690)
Admissions Rank	---	---	-0.0021 (0.0027)	---	---	0.0003 (0.0011)
Masters in Econ/Fin	---	---	-0.1388 (0.1580)	---	---	0.0140 (0.0618)
Pseudo R2	0.01	0.06	0.06	0.02	0.07	0.05
P-value for GRE Scores	0.402	0.605	0.124	0.174	0.795	0.452
Observations	871	839	468	884	852	468

Notes: Job types are divided into 4 categories: 1 = independent research; 2 = teaching; 3= dependent research; 4=business sector and noncompleted Ph.D. Columns 1-3 are estimated by ordered probit and columns 4-6 are estimated by binary probit model. P-value for GRE scores is from an F-test of the null hypothesis that the 3 scores jointly equal zero.

* = p-value < .10 **=p-value<.05